# FIN 620 Emp. Methods in Finance

#### Lecture 7 – Natural Experiment [P2]

Professor Todd Gormley

#### Announcements

- Upload rough draft of proposal to Canvas by noon today
  - I will try to get these graded quickly so that you have my feedback in time to adjust before final proposal
- Because of a conflict, our next class will be on Friday, 1-4pm in KH 301

# Informal Survey

#### Please fill out informal survey

https://forms.gle/jxgSeRhpxoWQ4qY2A

- Helps me figure out what changes I can make to improve the course for the second half and for future students
  - E.g., what topic should I have spent more time on? What topic did you find the most interesting? Is there too much, or too little work? Etc.

# Background readings

Roberts and Whited
Sections 2.2, 4

Angrist and Pischke
 Section 5.2

# Outline for Today

- Quick review of last lecture
- Continue to discuss natural experiments
  - □ How to handle multiple events
  - Triple differences
  - Common robustness tests that can be used to test whether internal validity is likely to hold
- Student presentations of "NE #1" papers

# Quick Review[Part 1]

- Natural experiment provides random variation in x that allows causal inference
  - Can be used in IV, regression discontinuity, but most often associated with "treatment" effects
- Two types of simple differences
  - Post-treatment comparison of treated & untreated
    Pre- and post-treatment comparison of treated

# Quick Review [Part 2]

Difference-in-differences is estimated with...

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + u_{i,t}$$

- Compares <u>change</u> in y pre- versus post-treatment for treated to <u>change</u> in y for untreated
- Requires "parallel trends" assumption
- Let's test your ability to identify a violation of the necessary assumptions for simple diffs and diff-in-diffs...

# Quick Review [Part 3]

- Suppose Spain exits the Euro and Ann compares profitability of Spanish firms after the exit to profitability before...
- What is necessary for the comparison to have any causal interpretation?
  - Answer = We must assume profitability after
     Spain's exit would have been same as profitability
     prior to exit absent exit... Highly implausible

# Quick Review [Part 4]

- Now, suppose Bob compares profitability of Spanish firms after the exit to profitability of German firms after exit...
- What is necessary for the comparison to have any causal interpretation?
  - Answer = We must assume profitability of Spanish firm would have been same as profitability of German firms absent exit... Again, this is highly implausible

# Quick Review [Part 5]

- Lastly, suppose Charlie compares <u>change</u> in profitability of Spanish firms after exit to <u>change</u> in profitability of German firms
- What is necessary for the comparison to have any causal interpretation?
  - Answer = We must assume change in profitability of Spanish firm would have been same as change for German firms absent exit...
     I.e., parallel trends assumption

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
  - Using group means to get an estimate
  - □ When additional controls are appropriate
- How to handle multiple events
- Falsification tests
- Triple differences

#### Standard Regression Format

Difference-in-differences estimator

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + u_{i,t}$$

- $p_t = 1$  if period *t* occurs after treatment and equals zero otherwise
- $d_i = 1$  if unit is in treated group and equals zero otherwise
- But there is another way that just involves comparing four sample means...

#### Comparing group means approach

To see how we can get the same estimate,
 β<sub>3</sub>, by just comparing sample means, first calculate expected *y* under four possible combinations of *p* and *d* indicators

Comparing group means approach [P1]

Again, the regression is...

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + u_{i,t}$$

• And the four possible combinations are:

 $E(y | d = 1, p = 1) = \beta_0 + \beta_1 + \beta_2 + \beta_3$  What assumption did I  $E(y | d = 1, p = 0) = \beta_0 + \beta_2$  make in doing this?  $E(y | d = 0, p = 1) = \beta_0 + \beta_1$  Answer: E(u | d, p) = 0; i.e.,  $E(y | d = 0, p = 0) = \beta_0$  the "experiment" is random Comparing group means approach [P2]

$$E(y | d = 1, p = 1) = \beta_0 + \beta_1 + \beta_2 + \beta_3$$
  

$$E(y | d = 1, p = 0) = \beta_0 + \beta_2$$
  

$$E(y | d = 0, p = 1) = \beta_0 + \beta_1$$
  

$$E(y | d = 0, p = 0) = \beta_0$$

These can be arranged in two-by-two table

	Post-Treatment,	Pre-Treatment,	
	(1)	(2)	
Treatment, (a)	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_0 + \beta_2$	
Control, (b)	$\beta_0 + \beta_1$	β <sub>o</sub>	

Comparing group means approach [P3]

Now take the simple differences

	Post-Treatment, (1)	Pre-Treatment, (2)	Difference, (1)-(2)
Treatment, (a)	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_0 + \beta_2$	$\beta_1 + \beta_3$
Control, (b)	$\beta_0 + \beta_1$	β <sub>0</sub>	β <sub>1</sub>
Difference, (a)-(b)	$\beta_2 + \beta_3$	β <sub>2</sub>	

Comparing group means approach [P4]

#### Then, take difference-in-differences!

<b>Control</b> , (b) $\beta_0^-$	+β <sub>2</sub> +β <sub>3</sub> +β <sub>1</sub>	$\beta_0 + \beta_2$ $\beta_0$	$\beta_1 + \beta_3 \\ \beta_1$
<b>Difference</b> , (a)-(b) $\beta_2$	+β <sub>3</sub>	β <sub>2</sub>	β3

This is why they call it the difference-in-differences estimate; regression gives you same estimate as if you took differences in the group averages

# Simple difference – Revisited [Part 1]

#### Useful to look at simple differences

	Post-Treatment, (1)	Pre-Treatment, (2)	Difference, (1)-(2)
Treatment, (a) Control, (b)	$\beta_0 + \beta_1 + \beta_2 + \beta_3$ $\beta_0 + \beta_1$	$\beta_0 + \beta_2$ $\beta_0$	$eta_1 + eta_3 \ eta_1$
Difference, (a)-(b)	$\beta_2 + \beta_3$	β <sub>2</sub>	β <sub>3</sub>
This was cross-sectional simple difference			bes that simple diff of treatment, $\beta_3$ ?
		<b>Answer =</b> when $\beta_2$ equals zero; i.e., no difference in level of <i>y</i> absent treatment	

# Simple difference – Revisited [Part 2]

Now, look at time-series simple diff...

	Post-Treatment, (1)	Pre-Treatment, (2)	Difference, (1)-(2)
Treatment, (a)	$\beta_0 + \beta_1 + \beta_2 + \beta_3$	$\beta_0 + \beta_2$	$\beta_1 + \beta_3$
Control, (b)	$\beta_0 + \beta_1$	β <sub>0</sub>	β1
Difference, (a)-(b)	$\beta_2 + \beta_3$	β <sub>2</sub>	β <sub>3</sub>
This was time simple diffe		give effec Answer = whe	es that simple diff et of treatment, $\beta_3$ ? en $\beta_1$ equals zero; i.e., <i>y</i> absent treatment

#### Why the regression is helpful

- Some papers will just report this simple two-by-two table as their estimate
- But there are advantages to the regression
  - Can modify it to test timing of treatment [we will talk about this in robustness section]
  - $\Box$  Can add additional controls, X

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
  - Using group means to get an estimate
  - When additional controls are appropriate
- How to handle multiple events
- Falsification tests
- Triple differences

#### Adding controls to diff-in-diff

Easy to add controls to regression

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + \Gamma X_{i,t} + u_{i,t}$$

X is some vector of controls
 L is waster of coefficients

 $\Box \ \Gamma \text{ is vector of coefficients}$ 

 E[y | d,p] in prior proofs just becomes E[y | d,p,X]

From earlier lecture, what type of controls should you NEVER add?

#### When controls are inappropriate

- Remember! You should never add controls that might themselves be affected by treatment
  - □ Angrist-Pischke call this a "bad control"
  - You won't be able to get a consistent estimate of  $\beta_3$  from estimating the equation

# A Pet Peeve of TG – Refined

- If you have a treatment that is truly random, do not put in controls affected by the treatment!
  - □ I've had many referees force me to add controls that are likely to be affected by the treatment...
  - If this happens to you, put in both regressions (with and without controls), and at a minimum, add a caveat as to why adding controls is inappropriate

#### When controls are appropriate

Two main reasons to add controls

Improve precision (i.e., lower standard errors)
 Restore 'random' assignment of treatment

# #1 – To improve precision

- Adding controls can soak up some of residual variation (i.e., noise) allowing you to better isolate the treatment effect
  - □ Should the controls change the estimate?
    - NO! If treatment is truly random, adding controls shouldn't affect actual estimate; they should only help lower the standard errors!
  - If adding controls changes estimates, you might have 'bad controls' *or* worse, nonrandom treatment <sup>(S)</sup>

# *Example* – Improving precision

- Suppose you have firm-level panel data
- Some natural experiment 'treats' some firms but not other firms
  - Could just estimate the standard diff-in-diff

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + u_{i,t}$$

 Or, could add fixed effects (like firm and year FE) to get more precise estimate... *Example* – Improving precision [Part 2]

So, suppose you estimate...

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 (d_i \times p_t) + \alpha_i + \delta_t + u_{i,t}$$
  
Firm fixed effects Year fixed effects effects

What meaning does β<sub>1</sub> have now?
What meaning does β<sub>2</sub> have now?

# *Example* – Improving precision [Part 3]

- Trick question! They have no meaning!
  - *p<sub>t</sub>* is perfectly collinear with year FE [because it doesn't vary across firms]
  - *d<sub>i</sub>* is perfectly collinear with firm FE
     [because it doesn't vary across time for each firm]
- Stata just randomly drops a couple of the FE
  - The estimates on *p<sub>t</sub>* and *d<sub>i</sub>* are just random intercepts with <u>no</u> meaning

*Example* – Improving precision [Part 4]

Instead, you should estimate...

$$y_{i,t} = \beta_0 + \beta_3 (d_i \times p_t) + \alpha_i + \delta_t + u_{i,t}$$

Firm fixed effects control for treatment

Year fixed effects control for posttreatment

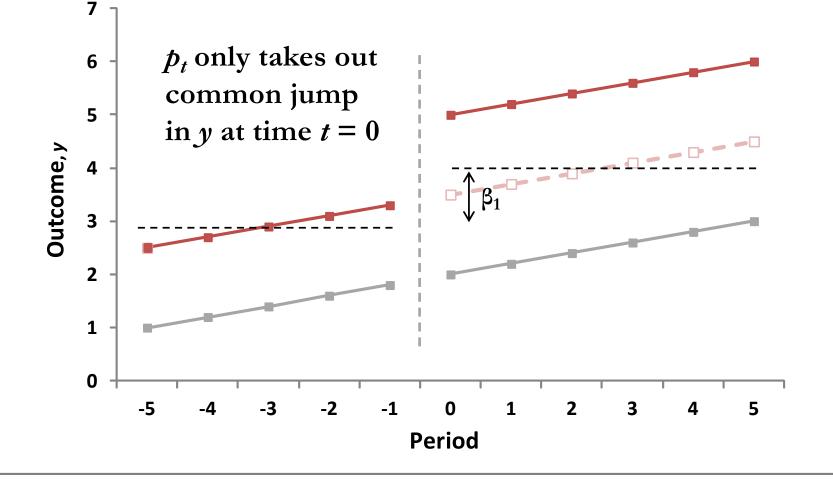
This is what some call the generalized difference-in-differences estimator

#### Generalized Difference-in-differences

- Advantage of generalized differences-indifferences is that it can improve precision and provide better fit of model
  - It doesn't assume all firms in treatment (or untreated) group have same average *y*; it allows intercept to vary for each firm
  - It doesn't assume that common change in y around event is a simple change in level; it allows common change in y to vary by year

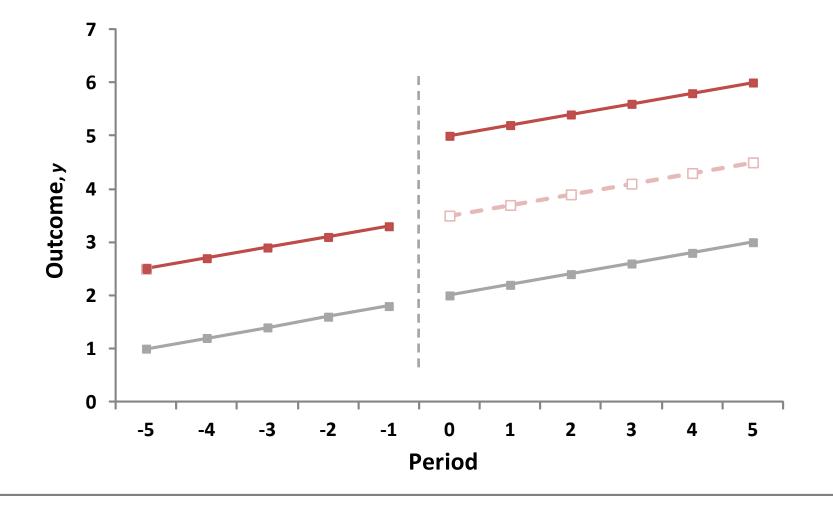
#### Generalized D-i-D – Example [Part 1]

 To see how Generalized D-i-D can be helpful, consider the example from last class



#### Generalized D-i-D – Example [Part 2]

Year dummies will better fit actual trend



#### When controls are appropriate

Two main reasons to add controls
Improve precision (i.e., lower standard errors)
Restore 'random' assignment of treatment

#### #2 – Restore randomness of treatment

- Suppose the following is true...
  - Observations of certain characteristic, 
     e.g., high x, are more likely to be treated
  - And firms with this characteristic are likely to have differential trend in outcome y
- Adding control for x could restore 'randomness'; i.e., being treated is random after controlling for x!

And nonrandomness is problematic for identification

I.e., treatment

isn't random

#### Restoring randomness – Example

- Natural experiment is change in regulation
  - □ Firms affected by regulation is random, except that it is more likely to hit firms that are larger
  - And, we think larger firms might have different trend in outcome *y* afterwards for other reasons
  - <u>And</u> firm size is not going to be affected by the change in regulation in any way
- If all true, adding size as control would be an appropriate and desirable thing to do

#### Controls continued...

- In prior example, suppose size is potentially affected by the change in regulation...
  - What would be another approach that won't run afoul of the 'bad control' problem?
    - Answer: Use firm size in year <u>prior</u> to treatment <u>and</u> its interaction with post-treatment dummy
    - This will control for non-random assignment (based on size) and differential trend (based on size)

Restoring randomness – Caution!

- In practice, don't often see use of controls to restore randomness
  - Requires assumption that non-random assignment isn't also correlated with <u>unobservable</u> variables...
  - So, not that plausible unless there are very specific reasons for non-randomness
- But regression discontinuity is one example of this; we'll see it next week

#### One last note... be careful about SEs

- Again, if have multiple pre- and post-treatment periods, need to be careful with standard errors
  - Either cluster SEs at level of each unit
  - Or collapse data down to one pre- and one posttreatment observation for each cross-section
- We will discuss more about standard errors in lecture on "standard errors"

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
- How to handle multiple events
  - Why they are useful
  - □ Simple estimation approach & its problems
  - Better ways to handle multiple events
- Falsification tests
- Triple differences

Motivating example...

- Gormley and Matsa (2011) looked at firms' responses to increased left-tail risk
  - Used discovery that workers were exposed to harmful chemical as exogenous increase in risk
  - One discovery occurred in 2000; a chemical heavily used by firms producing semiconductors was found to be harmful
- Can you think of any concerns about parallel trends assumption of this setting?

## Motivating Example – Answer

- Answer: Yes... This coincides with bursting of technology bubble; technology firms might arguably trend differently after 2000 for this reasons unrelated to chemical
  - How might multiple treatment events, occurring at different times (which is what Gormley and Matsa used), help?

#### Multiple treatment events

- Sometimes, the natural experiment is repeated a multiple points in times for multiple groups of observations
  - E.g., U.S. states make a particular regulatory change at different points in time
- These settings are particularly useful in mitigating concerns about violation of parallel trends assumption...

#### How multiple events are helpful

- Can show that effect of treatment is similar across different time periods
- Can show effect of treatment isn't driven by a particular set of treated firms
  - I.e., now the "identification police" would need to produce story as to why parallel trends is violated for <u>each</u> unique event

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
- How to handle multiple events
  - Why they are useful
  - □ Simple estimation approach & its problems
  - Better ways to handle multiple events
- Falsification tests
- Triple differences

#### Estimation with Multiple Events

- Estimating model with multiple events is still relatively easy to do
  - One can use approach of Bertrand and Mullainathan (JPE 2003)...

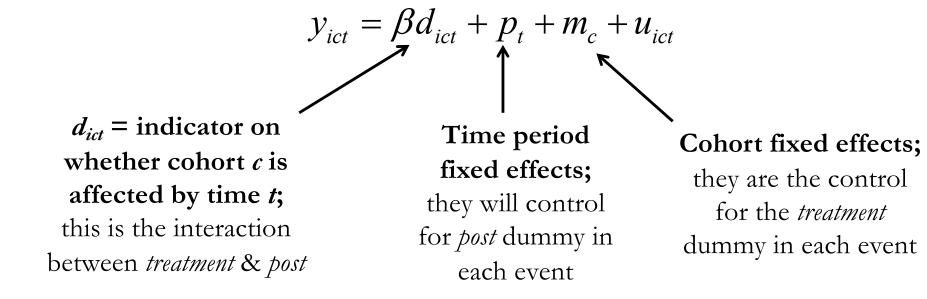
# Multiple Events [P1]

Just estimate the following estimation

$$y_{ict} = \beta d_{ict} + p_t + m_c + u_{ict}$$

- *y<sub>ict</sub>* is outcome for unit *i* (e.g., firm) in period *t* (e.g., year) and cohort *c*, where "cohort" indexes the different sets of firms <u>treated</u> by each event
  - E.g., different firms might be affected by a change in regulation at different points in time; firms <u>affected</u> at one point in time are a 'cohort'

# Multiple Events [P2]



# Multiple Events [P3]

- Intuition of this approach...
  - Every untreated observation at a particular point in time acts as control for treated observations in that time period
    - E.g., a firm treated in 1999 by some event will act as a control for a firm treated in 1994 until itself becomes treated in 1999
  - β will capture <u>average</u> treatment effect across the multiple events

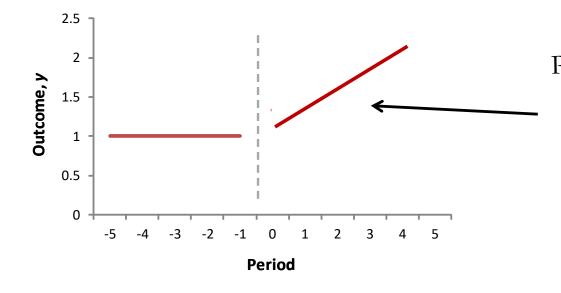
### However, a big potential problem...

- Early treated units are used as a control for later treated units
  - E.g., because they have already been treated, they don't change treatment status for the later units

Do you see a potential problem with this?

# A big potential problem... [Part 2]

Yes, there <u>will be problem</u> if treatment effect is dynamic, which can lead to violation of parallel trends! E.g., suppose treatment effect looks like...



Previously treated controls will *still be reacting* to treatment when used as later controls, violating parallel trends

# A big potential problem... [Part 3]

- Typical pre-trend test will also be biased, making it unhelpful in detecting problem
- See Baker, Larcker, and Wang (JFE 2022) for a nice description of the problem

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
- How to handle multiple events
  - Why they are useful
  - □ Simple estimation approach & its problems
  - Better ways to handle multiple events
- Falsification tests
- Triple differences

#### Multiple Events – <u>A better way</u>

- An alternative (and better way) to do a diffin-diffs using multiple events is to use the "stacked regression" approach developed by Gormley and Matsa (RFS 2011)
  - This approach will avoid bias from dynamic treatment effects and offers other advantages...

#### Stacked regression approach

 Now, think of running generalized diff-indiffs for just one of the multiple events...

$$y_{it} = \beta (d_i \times p_t) + \alpha_i + \delta_t + u_{it}$$

- *d<sub>i</sub>* = indicator for unit *i* (e.g., firm) being a treated firm in that event
- $p_t$  = indicator for treatment having occurred by period t (e.g., year)
- Unit *i* and period *t* FE control for the independent effects of  $d_i$  and  $p_t$

# Stacked regression approach [P2]

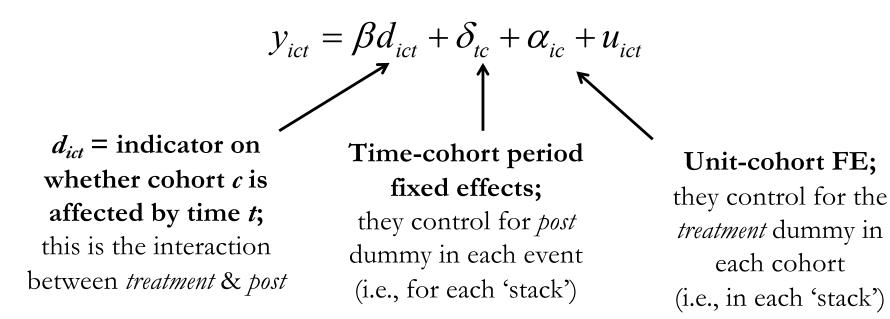
- But, contrary to standard difference-indifferences, your sample is...
  - Restricted to a small window around event;
     e.g., 5 years pre- and post- event
  - And drops any observations that are treated by <u>another</u> event
    - I.e., your sample starts only with previously untreated observations, and if a 'control' observation later gets treated by a different event, those <u>post</u>-event observations are dropped

Stacked regression approach [P3]

- Now, create a similar sample for <u>each</u> "event" being analyzed
- Then, "stack" the samples into one dataset and create a variable that identifies the event (i.e., 'cohort') each observation belongs to
  - Note: some observation units will appear multiple times in the data [e.g., firm 123 might be a control in event year 1999 but a treated firm in a later event in 2005]

Stacked regression approach [P4]

Then, estimate the following on the stacked dataset you've created



# Why stacking approach is better...

- Same intuition and approach of standard DiD, <u>but</u> has several advantages
  - Not subject to earlier bias from dynamic effects [removes previously treated firms as controls]
  - Can more easily isolate a particular window of interest around each event [instead of using all preand post observations as in other approach]
  - Can more easily extend this into a tripledifference type specification [more on that later]

#### Multiple Events – Other methods

- Calloway and Sant'Anna (JoE 2021)
- Sun and Abram (JoE 2021)
  - Both approaches avoid the dynamic problem by also estimating individual treatment effects of each event and then aggregate them to overall effect
  - But they lack some advantages (e.g., ability to do a triple-difference estimation) and the simplicity & flexibility of Gormley and Matsa (RFS 2011)

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
- How to handle multiple events
- Falsification tests
- Triple differences

#### Falsification Tests for DiD

- Can never directly test underlying identification assumption, but can do some falsification tests to support its validity
  - #1 Compare pre-treatment observables
  - #2 Check that timing of observed change in *y* coincides with timing of event *[i.e., no pre-trend]*
  - #3 Check for treatment reversal
  - #4 Check variables that shouldn't be affected
  - #5 Add a triple-difference

## #1 – Pre-treatment comparison [Part 1]

- Idea is that experiment 'randomly' treats some subset of observations
  - If true, then ex-ante characteristics of 'treated' observations should be like ex-ante characteristics of 'untreated' observations
  - Showing treated and untreated observations are comparable in dimensions thought to affect y can help ensure assignment was random

# #1 – Pre-treatment comparison [Part 2]

- If find ex-ante difference in some variable z, is difference-in-difference is invalid?
  - □ **Answer** = Not necessarily.
    - We need some story as to why units are expected to have differential trend in *y* after treatment (for reasons unrelated to treatment) that is correlated with *z* for this to be a problem for identification
    - And, even with this story, we could just control for *z* and its interaction with time
    - But what would be the lingering concern?

## #1 – Pre-treatment comparison [Part 3]

#### Answer = unobservables!

 If the treated and control differ ex-ante in observable ways, we worry they might differ in unobservable ways that related to some violation of the parallel trends assumption

# #2 – Check for pre-trend [Part 1]

- Like last lecture, can just allow effect of treatment to vary by period to nonparametrically map out the timing
  - "Parallel trends" suggest we shouldn't observe any differential trend prior to treatment for the observations that are eventually treated

#### #2 – Check for pre-trend [Part 2]

Estimate the following:

$$y_{i,t} = \beta_0 + \beta_1 d_i + \beta_2 p_t + \sum_t \gamma_t \left( d_i \times \lambda_t \right) + u_{i,t}$$

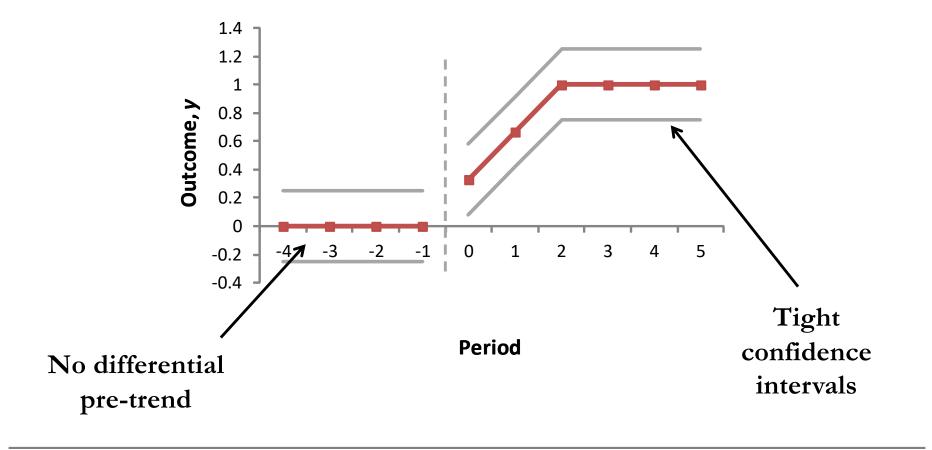
- *d<sub>i</sub>* and *p<sub>t</sub>* are defined just as before
   λ<sub>t</sub> is indicator that equals 1 if event time = t and zero otherwise, where
  - t = 0 is the period treatment occurs
  - t = -1 is period before treatment

# #2 – Check for pre-trend [Part 3]

- γ<sub>t</sub> estimates change in *y* relative to excluded periods; you then plot these in graph
  - Easiest to <u>fully saturate</u> the model (i.e., include λ<sub>t</sub> for every period but the very first one); then all estimates γ<sub>t</sub> are relative to this period
  - Can also plot confidence interval for each  $\gamma_t$

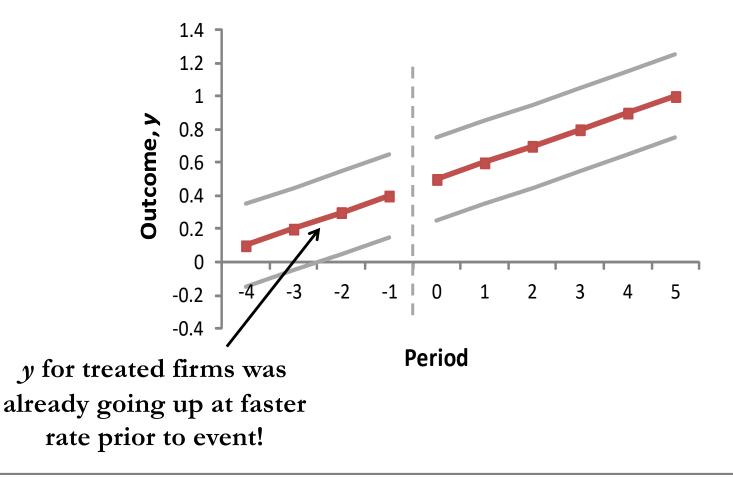
#2 – Check for pre-trend [Part 4]

Something like this is ideal...



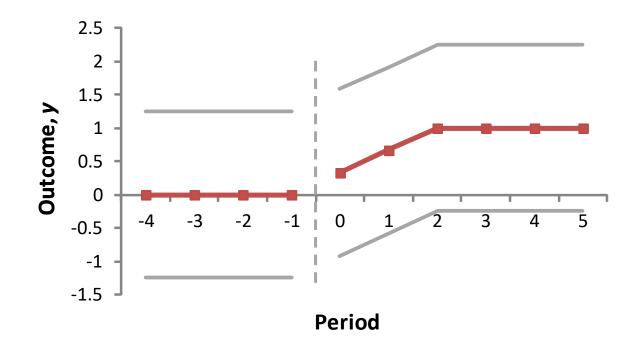
## #2 – Check for pre-trend [Part 5]

Something like this is very bad



#### #2 – Check for pre-trend [Part 6]

 Should we make much of wide confidence intervals in these graphs? E.g.



Answer: Not too much... Each period point estimate might be noisy; diff-in-diffs will tell us whether postaverage y is significantly different then preaverage y

# #2 – Check for pre-trend [Part 7]

- Another type of pre-trend check done is to do the diff-in-diffs in some "random" pretreatment to show no effect
  - □ I'm not a big fan of this... Why?
    - Answer #1 It is subject to gaming; researcher might choose a particular pre-period to look at that works
    - Answer #2 Prior approach allows us to see what the timing was and determine whether it is plausible

### #3 – Treatment reversal

- In some cases, the "natural experiment" is subsequently reversed
  - E.g., regulation is subsequently undone
- If we expect the reversal should have the opposite effect, it is good to confirm this

#### #4 – Unaffected variables

- In some cases, theory provides guidance on what variables should be unaffected by the "natural experiment"
  - □ If natural experiment is what we think it is, we should see this in the data... so check

## #5 – Add Triple difference

 If theory tells us treatment effect should be larger for one subset of observations, we can check this with triple difference

□ Pre- *versus* post-treatment

Untreated versus treated

Less sensitive *versus* more sensitive
 [or larger *versus* smaller treatment level]

This is the third difference

# Natural Experiment Outline – Part 2

- Difference-in-difference continued...
- How to handle multiple events
- Falsification tests
- Triple differences
  - □ How to estimate & interpret it
  - Using the popular subsample approach

#### Diff-in-diffs-in-diffs – Regression

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 h_i + \beta_4 (p_t \times h_i) + \beta_5 (d_i \times h_i) + \beta_6 (p_t \times d_i) + \beta_7 (p_t \times d_i \times h_i) + u_{i,t}$$

- $p_t = 1$  if period *t* occurs after treatment and equals zero otherwise
- $d_i = 1$  if unit is in treated group and equals zero otherwise
- *b<sub>i</sub>* = 1 if unit is group that is expected to be more sensitive to treatment [or experience a larger treatment level]

## Diff-in-diff-in-diff – Regression [Part 2]

- How to choose and set  $b_i$ 
  - E.g., If theory says effect is bigger for larger firms; could set  $h_i = 1$  if assets of firm in year prior to treatment is above the median size
  - Note: Remember to use <u>ex-ante</u> measures to construct indicator if you think underlying variable (that determines sensitivity) might be affected by treatment... Why?

**Answer =** To avoid bad controls!

#### Diff-in-diff-in-diff – Regression [Part 3]

$$y_{i,t} = \beta_0 + \beta_1 p_t + \beta_2 d_i + \beta_3 h_i + \beta_4 (p_t \times h_i) + \beta_5 (d_i \times h_i) + \beta_6 (p_t \times d_i) + \beta_7 (p_t \times d_i \times h_i) + u_{i,t}$$

- Easy way to check if done correctly...
  - Should have 8 coefficients (including constant) to capture the 2×2×2=8 different combinations
  - Likewise, a double difference has 4 coefficients (including constant) for the 2 × 2=4 combinations
- What do  $\beta_6$  and  $\beta_7$  capture?

Interpreting the estimates [Part 1]

- β<sub>6</sub> diff-in-diff estimate
   for the less-sensitive obs.
  - Captures average <u>differential</u> change in *y* from the pre- to post-treatment period for the less sensitive observations in the treatment group
     *relative* to the change in *y* for the less sensitive observations in the untreated group

## Interpreting the estimates [Part 2]

- β<sub>7</sub> is the triple diff estimate; it tells us how much larger effect is for the more sensitive obs.
  - β<sub>7</sub> captures how different the difference-indifferences estimate is for observations considered more sensitive to the treatment [or observations that receive a larger treatment level]
  - What is total treatment effect for these firms?
  - $\Box \text{ Answer} = \beta_6 + \beta_7$

## Tangent – Continuous vs. Indicator?

- Can also do the triple difference replacing *h<sub>i</sub>* with a continuous measure instead of indicator
  - E.g., suppose we expect treatment effect is bigger for larger firms; rather than constructing indicator based on ex-ante size, could just use ex-ante size
  - □ What are the advantages, disadvantages of this?

## Tangent – Continuous vs. Indicator?

#### Advantages

- Makes better use of variation available in data
  Provides estimate on magnitude of sensitivity
- Disadvantages
  - Makes linear functional form assumption; indicator imposes less structure on the data
  - □ More easily influenced by outliers

## Generalized Triple-Difference

- Like diff-in-diffs, can add in FE to soak up the various terms and improve precision
- E.g., in firm-level panel regression with firm and year fixed effects, you'd estimate

$$y_{i,t} = \beta_1 (p_t \times h_i) + \beta_2 (p_t \times d_i)$$
  
+  $\beta_3 (p_t \times d_i \times h_i) + \delta_t + \alpha_i + u_{i,t}$ 

The other terms (including the constant) all drop out; they are collinear with the FE

# Natural Experiment [P2] – Outline

- Difference-in-difference continued...
- How to handle multiple events
- Falsification tests
- Triple differences
  - □ How to estimate & interpret it
  - Using the popular subsample approach

# Subsample Approach

- Instead of doing full-blown triple-difference, you can also just estimate the double-difference in the two separate subsamples
  - Double-difference for low sensitive obs. (i.e., *h<sub>i</sub>* = 0)
     Double-difference for more sensitive obs. (i.e., *h<sub>i</sub>* = 1)
- Note: the estimates won't directly match the  $\beta_2$ ,  $\beta_2 + \beta_3$  effects in prior estimation... Why?

# Subsample Approach Differences...

- Answer = In subsample approach year FE are allowed to differ by sub-sample
  - Therefore, subsample approach is controlling for more things
  - However, one can easily recover the subsample estimates in one regression (and test the statistical difference) between subsamples by estimating...

# Matching Subsample to Combined [P1]

$$y_{i,t} = \beta_2 \left( p_t \times d_i \right) + \beta_3 \left( p_t \times d_i \times h_i \right) + \delta_t + \left( \delta_t \times h_i \right) + \alpha_i + u_{i,t}$$

Year FE interacted with sensitivity indicator

- Just add interaction between year FE and indicator for being more sensitivity...
  - This allows for different year FE for each subsample, which is what happened when we estimated the subsamples in two separate regressions

# Matching Subsample to Combined [P2]

#### In prior regression...

- β<sub>2</sub> will equal coefficient from diff-in-diffs using just the subsample of less sensitive observations
- □  $\beta_2 + \beta_3$  will equal coefficient from diff-in-diffs using just the subsample of more sensitive observations
- *t*-test on β<sub>3</sub> tells you whether effect for more sensitive subsample is statistically different from that of the less sensitive subsample

# Triple Diff – Stacked Regression [Part 1]

- Another advantage of <u>stacked</u> regression approach to multiple events is ability to more easily incorporate a triple diff
  - Can simply run stacked regression in separate subsamples to create triple-diffs or run it in one regression as shown previously

# Triple Diff – Stacked Regression [Part 2]

- Can't easily do either of these in approach of Bertrand and Mullainathan (2003)
  - Some observations act as both 'control' and 'treated' at different points in sample; not clear how create subsamples in such a setting
  - With Gormley and Matsa (2011) stacked approach, however, you can create subsample for each stack based on characteristic of treated and control firms in year prior to treatment

## External Validity – Final Note

- While randomization ensures internal validity (i.e., causal inferences), external validity might still be an issue
  - Is the experimental setting representative of other settings of interest to researchers?
    - I.e., can we extrapolate the finding to other settings?
    - A careful argument that the setting isn't unique or that the underlying theory (for why you observe what you observe) is likely to apply elsewhere is <u>necessary</u>

# Summary of Today [Part 1]

- Diff-in-diffs & control variables
  - Don't add controls affected by treatment
  - Controls shouldn't affect estimates, but can help improve precision
- Multiple events are helpful in mitigating concerns about parallel trends assumption
  - But follow Gormley and Matsa (2011) to avoid potential bias from dynamic treatment effects

# Summary of Today [Part 2]

- Many falsification tests one should do to help assess internal validity
  - Ex. #1 Compare ex-ante characteristics
    Ex. #2 Check timing of observed effect
- Triple difference is yet another way to check internal validity and mitigate concerns about identification

## In First Half of Next Class

- Regression discontinuity
  - □ What are they?
  - □ How are they useful?
  - How do we implement them?
- Related readings... see syllabus

Assign papers for next week...

Gormley and Matsa (RFS 2011)

□ Risk & CEO agency conflicts

- Becker and Stromberg (RFS 2012)
  - □ Agency conflicts between equity & debt
- Ashwini (JFE 2012)
  - □ Investor protection laws & corporate policies

## Break Time

- Let's take our 10-minute break
- We'll do presentations when we get back